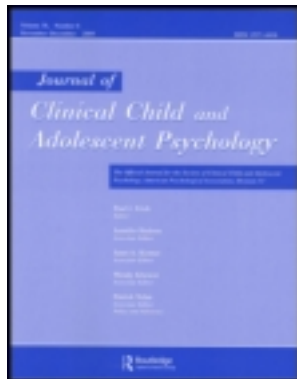


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Interaction of 5-HTTLPR and Idiographic Stressors Predicts Prospective Depressive Symptoms Specifically Among Youth in a Multiwave Design

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5-HTTLPR, episodic stressors, depressive and anxious symptoms were assessed prospectively (child and parent report) every 3 months over 1 year (5 waves of data) among community youth ages 9 to 15 ($n = 220$). Lagged hierarchical linear modeling analyses showed 5-HTTLPR interacted with idiographic stressors (increases relative to the child's own average level over time), but not nomothetic stressors (higher stress exposure relative to the sample), to predict prospective elevations in depressive, but not anxious, symptoms. Youth with copies of the S or L_G alleles of 5-HTTLPR, who experienced more stressors relative to their typical level, exhibited prospective increases in depressive symptoms over time. These findings suggest that 5-HTTLPR confers susceptibility to depression via stress reactivity.

Most individuals experience their first depressive episode in adolescence (Costello Mustillo, Erkanli, Keeler, & Angold, 2003), and adolescent-onset depression substantially increases risk for continuity into adulthood (Rutter, Kim-Cohen, & Maughan, 2006). Investigating gene–environment interactions offers a promising window to study putative pathophysiological and psychosocial risk mechanisms (Moffitt, Caspi, & Rutter, 2006). Caspi and colleagues (2003) demonstrated that a measured genotype (5-HTTLPR, the serotonin transporter-linked polymorphic region) by environment (major negative events) interaction ($G \times E$) predicted

increases in depression in adults. However, since this seminal paper, findings testing $G \times E$ with 5-HTTLPR and various environmental risks among adults have been mixed (Monroe & Reid, 2008; Risch et al., 2009).

Focus on 5-HTTLPR as a candidate gene of interest in the pathophysiology of depression arose after publications implicated the polymorphism in several biological phenotypes of the disorder including neuroticism (Lesch et al., 1996) and amygdala reactivity to threatening stimuli (Hariri et al., 2002). Despite initial support for 5-HTTLPR interacting with negative life events to predict depression in adults (Caspi et al., 2003), subsequent findings have been mixed, with some studies demonstrating a clear $G \times E$ (e.g., Caspi et al., 2003; Taylor et al., 2006; Wilhelm et al., 2006; Zalsman et al., 2006) and others only partially replicating or failing to replicate these results (e.g., Grabe et al., 2005; Gillespie, Whitfield, Williams, Heath, & Martin, 2005; Kendler, Kuhn, Vittum, Prescott, & Riley, 2005).

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Recently, Risch and colleagues (2009) published a meta-analysis that sought to include exact replications of the original Caspi and colleagues (2003) 5-HTTLPR × Stress interaction predicting depression. Results showed no overall significant G × E for depression but did reveal a significant main effect for stress. However, as several subsequent publications have noted, this meta-analysis is limited in numerous ways (Caspi, Hariri, Holmes, Uher, & Moffitt, 2010; Rutter, Thapar, & Pickles, 2009). The primary aim of the current study was to examine G × E effects on depressive symptoms, specifically among youth, using a powerful multiwave design. In addition, we sought to address some of the limitations of prior studies included in the Risch et al. meta-analysis, including measurement precision for genetic risk and stressor exposure as well as approaches to testing G × E (i.e., idiographic and nomothetic approaches to stress reactivity; see next) to create a study design with more power to detect important etiological factors in the pathophysiology of depression.

The vast majority of the G × E literature in depression has focused predominantly on adult samples. Studying gene–environment interactions among youth is a relatively understudied area of importance given the strong risk for recurrence of adolescent onset depression. Of the extensive research literature, based primarily on adult samples, there are a few studies that have focused on youth samples and found an interaction between 5-HTTLPR and environment, including maltreatment (Cicchetti, Rogosch, & Sturge-Apple, 2007, *M* age = 17; Kaufman et al., 2006, *M* age = 9), general stressors (Eley et al., 2004; 12–19 years old), family structure (Nobile et al., 2009; ages 10–14), and chronic family stress (Hammen, Brennan, Kennan-Miller, Hazel, & Najman, 2010; ages 15 and 20). Moreover, most of the studies with youth samples are limited by cross-sectional and mono-informant designs (see Hammen et al., 2010, as an exception).

Given mixed G × E evidence among adults and few, limited studies among youth, the current study examined G × E effects on depressive symptoms specifically in a multiwave study of youth. We examined the hypothesis that 5-HTTLPR allelic variation would moderate the longitudinal association between idiographic stress exposure (exposure relative to one's own average) and elevations in depressive symptoms specifically, compared to anxious symptoms, over time and explored whether this longitudinal G × E effect would be stronger among girls or vary by youths' age (as indexed by grade level).

Polymorphisms of the 5HT system are candidates of interest and biologically plausible for testing G × E effects on depression given 5HT's involvement in emotion and cognition (Canli & Lesch, 2007). The serotonin transporter regulates serotonin function by terminating

serotonin action in the synapse via reuptake. The short (S) allele is associated with decreased transcriptional efficiency compared with the long (L) allele (Canli & Lesch, 2007). The decreased transcriptional efficiency associated with the S allele results in less serotonin being recaptured in the presynaptic neuron when compared to the L allele. Although the exact mechanism by which this polymorphism gives rise to psychiatric outcomes has not been fully elucidated, there have been many studies investigating its role in depression (e.g., Brown & Harris, 2008; Gibb, Benas, Grassia, & McGeary, 2009; Gibb, Uhrlass, Grassia, Benas, & McGeary, 2009; Gotlib, Joormann, Minor, & Hallmayer, 2008; Hammen et al., 2010; Levinson, 2006; see recent review by Caspi et al., 2010). Findings suggest that although variations in 5-HTTLPR are not directly correlated with depression (Lesch, 2003), 5-HTTLPR may interact with other factors (e.g., attentional biases, negative cognitive styles, early adversity, negative life stress) to place individuals at higher risk for experiencing symptoms of depression. More recently, a common single nucleotide polymorphism (SNP) was found to occur (adenine to guanine) in the L allele (rs25531). Only the L_A allele is high functioning (referred to as L), whereas the L_G allele is more functionally equivalent to the S allele (collectively referred to as S; Hu et al., 2006; Hu et al., 2005). Referred to as the triallelic method, this SNP is considered a better reflection of the transcriptional efficiency of 5-HTTLPR (Hu et al., 2006). The biallelic S and L approach has been most reported in the literature; however, this approach does not separate the L_G from the L_A alleles, which could result in individuals being incorrectly classified as low risk. It has been recommended that both triallelic and biallelic genotyping approaches be reported (Martin, Cleak, Willis-Owen, Flint, & Shifman, 2007) and we used both approaches to genotyping 5-HTTLPR in the current study.

In addition to genotyping, a reliable and valid assessment of stressors is essential for testing and interpreting G × E. Yet, most of the stress measures used in the adult G × E studies have unknown psychometric properties (Monroe & Reid, 2008). It is important to note that comprehensive reviews of the G × E literature have found that methodological differences in stressor measurement affected whether studies replicated the original Caspi and colleagues finding (Caspi, Ahmad, Holmes, Uher, & Moffitt, 2010; Uher & McGuffin, 2010). Studies using more specific and clearly defined negative events and stress measures that are psychometrically strong were significantly more likely to find G × E effects in depression. Moreover, prior studies used assessments of stressors with variable time frames (e.g., 6 months in Eley et al., 2004, to 5 years in Caspi et al., 2003). Such imprecise timing of stressor occurrence

can be problematic for $G \times E$ in depression because most of the $G \times E$ studies, especially those included in the Risch et al. (2009) meta-analysis, used broad, distal measures of environmental risk, as opposed to more specific, proximal episodic stressors (e.g., within past 3 months) that have been shown temporally to maximally predict increases in depression (Monroe & Reid, 2008). In particular, it is unclear how measures of distal stress would affect serotonergic function and, in turn, risk to depression. Indeed, Strickland and colleagues (2002) suggested that stressors should affect serotonergic function within a short time frame following the stressor. As such, using a shorter time frame for stressor assessment may provide a stronger and more conceptually meaningful test of $G \times E$ in depression. In sum, to test $G \times E$ predicting depressive symptoms, it is as important to assess the environment (i.e., proximal stressors) in a reliable and developmentally valid manner as it is to ascertain genetic risk (Monroe & Reid, 2008).

Accurately assessing genetic and environmental risks to depression is essential. But equally vital is conceptually considering how genes interact with the environment to predict increases in depressive symptoms and incorporating a design that enables a rigorous test of hypothesized $G \times E$ processes. The current hypothesis is that 5-HTTLPR may affect individuals' sensitivity or reactivity to the environment (Caspi et al., 2010). Various lines of evidence, including human observational, neuroimaging, and animal research, is consistent with a stress reactivity conceptualization (see Caspi et al., 2010, for review). For example, Gotlib and colleagues (2008) showed that girls with the S allele of 5-HTTLPR reacted with greater cortisol release to a laboratory stressor. Experimental neuroscience research illustrates that S carriers exhibit exaggerated amygdala activity to threat-related stimuli (e.g., Hariri et al., 2002). Finally, animal research supports 5-HTTLPR as a gene that may affect stress reactivity (e.g., Li et al., 1999). To test the hypothesis that 5-HTTLPR may increase susceptibility to depression via stress reactivity, a multiwave design, in which stressors and symptoms are assessed repeatedly, can increase precision and power.

A multiwave design with repeated assessments of symptoms and stressors enables a more precise investigation of how stress exposure over time is related to elevations in depressive symptoms as a function of 5-HTTLPR allelic variation. When testing this stress reactivity hypothesis using multiwave data, one can distinguish between *nomothetic* versus *idiographic* stress conceptualizations and approaches to analysis (Abela & Hankin, 2008). The nomothetic approach states that individuals who experience high stress levels, compared to the *sample as a whole*, regardless of their individual fluctuations in stress over time, are the most likely to become depressed, especially when vulnerable. In

contrast, the idiographic approach suggests that individuals who experience higher stress exposure, compared to *their own average* over time and regardless of their rank-order in the sample, are most likely to become depressed, especially when vulnerable. To our knowledge, no prior $G \times E$ study has investigated stress reactivity as a function of genotype to predict depressive symptoms from nomothetic and idiographic perspectives. It is important to note, however, that research testing cognitive vulnerability interacting with stress, which is a stress reactivity hypothesis grounded in cognitive theories of depression (e.g., Abramson, Metalsky, & Alloy, 1989; Beck, 1987), has found stronger and more consistent support when using an idiographic as opposed to nomothetic approach when examining the impact of stress on depressive symptoms as a function of cognitive risk in adults (Abela, Webb, Ho, Wagner, & Adams, 2006; Gibb, Beevers, Andover, & Holleran, 2006; Stone, Gibb, & Coles, 2010) and youth (Abela, Skitch, Adams, & Hankin, 2006; Abela, Zuroff, Ho, Adams, & Hankin, 2006; see Abela & Hankin, 2008, for review). Testing an idiographic approach to stress requires multiwave data to ascertain individuals' average stress exposure and deviations from their own average (i.e., increases or decreases from typical stress exposure). However, there are very few multiwave $G \times E$ studies of depression (Gibb, Benas, et al., 2009, 8–12 years old; Gibb, Uhrlass, et al., 2009, 8–12 years old). To our knowledge, no study has examined how 5-HTTLPR interacts with stressors, conceptualized nomothetically as well as idiographically, to predict prospective changes in symptoms. All of the studies included in the Risch et al. (2009) meta-analysis conceptualized stress reactivity in a nomothetic manner.

Research has demonstrated that anxiety and depression are highly comorbid in youth (Chavira, Stein, Bailey, & Stein, 2004; Costello et al., 2003; Lewinsohn, Zinbarg, Seeley, Lewinsohn, & Sack, 1997). In addition, prospective studies have shown that anxiety symptoms and disorders tend to precede depression in youth (Beesdo et al., 2007; Pine, Cohen, Gurley, Brook, & Ma, 1998). Given the degree of comorbidity, as well as the temporal ordering of these two internalizing syndromes, it is important to investigate whether $G \times E$ is a shared etiology or specific to predicting depressive symptoms. The few studies that have examined the role of 5-HTTLPR interacting with stressful life events to predict anxiety in youth (Cicchetti et al., 2007; Gunthert et al., 2007; Olsson et al., 2005) and adults (Laucht et al., 2009; Stein, Schork, & Gelernter, 2008; Xie et al., 2009) have yielded mixed findings. In addition, both the type of anxiety examined (e.g., anxiety sensitivity, posttraumatic stress disorder, and state anxiety) as well as the nature of the environmental risk (e.g., childhood adversity, mild daily stressors, and

insecure attachment style) were highly variable among the studies. Concurrently examining both anxious and depressive symptoms among youth provides a stringent test of the specificity of $G \times E$ in their etiology.

In summary, this study examined gene–environment interactions, specifically the interaction of the 5-HTTLPR polymorphism with proximal episodic stressors over time, to predict prospective elevations in depressive symptoms among youth. Moreover, we investigated the specificity of this $G \times E$ to predict depressive versus anxious symptoms, which commonly co-occur with depression (Costello et al., 2003). Last, we explored the potential moderation of this $G \times E$ interaction by sex and age given prior research showing that the 5-HTTLPR \times Stress interaction was associated with depression among adolescent girls only (e.g., Eley et al., 2004; Hammen et al., 2010), behavioral genetic research finding $G \times E$ among adolescent girls (Silberg et al., 1999), and the sex difference in depression emerging in early adolescence and diverging strongly through middle adolescence (Hankin & Abramson, 2001).

To provide a more exacting test of $G \times E$, we followed Rutter's (2005) recommendations for methodologically rigorous developmental psychopathology research and incorporated (a) a multiwave longitudinal design allowing for more precise temporal ordering of events before depression, (b) a community sample of youth, and (c) multiple informants of key constructs such as depressive symptoms and stressors, which reduces common source information bias. In particular, we assessed developmentally salient episodic stressors and anxious and depressive symptoms every 3 months over 1 year (five waves of data) with both youth and parent report in a cohort design including samples of third-, sixth-, and ninth-grade youth. This powerful multiwave design enabled a more exacting examination of processes over time, that is not possible with cross-sectional or even two time point studies (Curran & Willoughby, 2003), and provided more statistical power (Snijders & Bosker, 1999) for testing $G \times E$ predictions.

We hypothesized that youth carrying at least one copy of the 5-HTTLPR S variant with less transcriptional efficiency, who encountered more idiographic stressors over time, would report greater prospective elevations specifically in depressive and not anxious symptoms over time. We did not expect to find a significant $G \times E$ predicting depression when stressors were analyzed in a nomothetic manner, consistent with the nonsignificant $G \times E$ effect observed in the Risch et al. (2009) meta-analysis using these methods, as an idiographic approach to stress reactivity has proved to be more consistent and strongly predictive of future depressive symptoms (Abela & Hankin, 2008). Finally, given previous findings of a $G \times E$ predicting depression among adolescent girls only (Eley et al., 2004; Hammen

et al., 2010) and the paucity of research investigating children, we examined whether sex and/or age moderates the $G \times E$ interaction in depressive symptoms.

METHOD

Participants

Children and adolescents were recruited by brief information letters sent home directly by the participating school districts to families with a child in third, sixth, and ninth grades of public schools. Approximately 2,000 families had a child in third, sixth, or ninth grade in a participating school district and therefore were eligible to receive letters. The short letter stated that we were conducting a study on social and emotional development in children and adolescents and requested that interested participants call the laboratory to receive more detailed information on the study. Of the brief introductory letters sent by the school districts, 305 families were counted as contacted, as a parent responded to the letter and called the laboratory for more information. Briefly, the 305 contacted parents called the laboratory and responded to a brief phone screen that established that both the parent and child were fluent in English, that the child did not carry an autism spectrum or psychotic disorder, and that the child had an IQ greater than 70. Of the 305 initially contacted families, 220 (72%) qualified as a study participant as they met criteria and arrived at the laboratory for the first assessment. The remaining 85 (28%) are considered non-participants for the following reasons: 3 (1%) were excluded due to child with autism spectrum disorder or non-English-speaking family, 12 (4%) declined after learning about the study's requirements, 49 (16%) were scheduled but did not arrive for assessment, and 21 (7%) were never reached to schedule an assessment. In sum, the present study has a participation rate of 72%. Our participation rate is above the 70% rate recommended by epidemiologists for having a representative sample of the target population (Tolonen, 2005; Tolonen, Dobson, & Kulathinal, 2005) and is comparable to that found in previous community-based, general samples of youth depression (e.g., 61% in Lewinsohn, Hops, Roberts, Seeley, & Andrews, 1993).

Participants were 220 youth ranging in age from 9 to 15 ($M = 11.4$, $SD = 2.27$). The sample was approximately evenly divided by sex (43% boys, 57% girls) and grade (31% third grade, 41% sixth grade, 28% ninth grade). The present sample, drawn from the general community of youth attending public schools, was representative of both the broader population of the particular geographical area and school districts from which the sample was drawn, including socioeconomic status,

ethnicity, and race. Ethnicity was as follows: Caucasian, 68%; African American, 22%; Latino, 5%; Asian/Pacific Islander, 3%; Other/Mixed ethnicity and race, 2%, which is comparable to that of the community and school districts from which the sample was recruited (Caucasian, 65%; African American, 11%; Latino, 19%; Asian/Pacific Islander, 4%; American Indian, 1%); 25% of youth from the entire school districts received free/reduced lunch. Sex and ethnicity did not differ and were approximately evenly distributed across grades (e.g., 44% boys in third, 46% in sixth, and 40% in ninth grades, respectively). Parents of the youth were predominantly mothers (87%), and 73% were married, 11% single, 15% divorced or separated, and 1% widowed. Median annual parental income was \$70,000 (range = \$10,000–\$200,000), and 24% of the youth received free/reduced lunch at school.

Procedures

The parent and youth visited the laboratory for the baseline assessment. Parents provided informed written consent for their participation and for their child; youth provided written assent. The initial baseline assessment consisted of a battery of questionnaires completed by youth and parents about the child and collection of youth DNA via saliva. Regular follow-up assessments over the phone with parent and youth occurred every 3 months over a 1-year period (five waves of data) to assess stressors and symptoms. The Institutional Review Board approved all procedures. Youth and parents were reimbursed for participation at baseline and each follow-up.

Measures

Depressive symptoms. The Children's Depression Inventory (CDI; Kovacs, 1985) assessed youths' depressive symptoms. Both the child (CDI-C) and parent (CDI-P) reported on the child's level of depressive symptoms to enable multiple informants of depression. CDI-C and CDI-P were given at all five assessments. CDI-C and CDI-P scores were moderately correlated (r_s ranged from .34 to .44, $p_s < .001$), so they were standardized and averaged together to form an overall depressive symptoms score at each wave. There extensive evidence to support the reliability and validity of the CDI as a measure of depressive symptoms in children (Klein, Dougherty, & Olino, 2005). In the current sample, internal consistency (α) was above .80 at all waves. At baseline, the range of CDI scores (child report: M CDI = 6.87, SD = 5.0, range = 0–35; parent report: M CDI = 5.69, SD = 5.6, range = 0–28) was comparable to published norms (Kovacs, 1985) and prior research with general community samples (Cole, Peeke, Martin, Truglio, & Seroczynski, 1998; Petersen

et al., 1993). Using recommended clinical cutoffs for the CDI (i.e., CDI scores >13; Kazdin, 1989; Smucker, Craighead, Craighead, & Green, 1986) revealed that 15.9% of youth report and 13.8% of parent report were above cut-scores for the CDI at baseline.

Anxious symptoms. The Multidimensional Anxiety Scale for Children (MASC; March, Sullivan, & Parker, 1999) assessed youths' overall anxious symptoms and was completed by both the child (MASC-C) and parent about the child (MASC-P). MASC-C and MASC-P were given at all five assessments. MASC-C and MASC-P scores were moderately correlated (r_s ranged from .18 to .31, $p_s < .001$) consistent with prior research (Baldwin & Dadds, 2007), so they were standardized and averaged together to form an overall anxious symptoms score at each wave. There is evidence to support the reliability and validity of the MASC as a measure of anxiety in children (Silverman & Ollendick, 2005). In the current sample, internal consistency (α) was above .80 at all waves. At baseline, the range of MASC scores (child report: M MASC = 43.26, SD = 22.2, range = 5–96; parent report: M MASC = 43.99, SD = 19.96, range = 15–91) was comparable to published norms (March et al., 1999). Using recommended clinical cutoffs for the MASC (March et al., 1999) revealed that 12% of youth report and 7% of parent report were above cut-scores (T scores >65) for the MASC at baseline.

Stressors. The Adolescent Life Events Questionnaire (ALEQ; Hankin & Abramson, 2002) consists of 37 items that assess the number of stressors occurring within the past 3 months. The ALEQ assesses a broad range of negative life events that typically occur among adolescents, including school, friendship, romantic, and family events. Respondents indicated whether the event occurred within the past 3 months, which resulted in a count of stressors over the last 3 months. Based on methodological recommendations on assessing stress and adversity (e.g., Brewin, Andrews, & Gotlib, 1993), the ALEQ includes concrete, specific anchors and descriptions of events rather than overly general subjective event descriptions commonly seen in other self-report stress checklists. The use of concrete anchors is intended to help reduce overreporting of trivial events or misinterpretation of events due to personality, mood, or memory bias. For example, instead of the event "did poorly on a test," which could be prone to misinterpretation (e.g., a straight-A student receives a B+ and endorses the event), the ALEQ specifically defines and anchors the event: "Did poorly on—i.e., C or worse, or failed, a test or class project." Both the child (ALEQ-C) and parent (ALEQ-P) reported on the child's exposure to stressors by indicating whether a stressor

occurred within the last 3 months. ALEQ-C and ALEQ-P were given at all five assessments. ALEQ-C and ALEQ-P scores were moderately correlated (r s ranged from .37 to .41, p s < .001), so they were standardized and averaged together to form an overall score at each wave. Scores for ALEQ-C ranged from 0 to 35 ($M = 16.5$, $SD = 7.5$), and ALEQ-P scores ranged from 2 to 37 ($M = 17.2$, $SD = 7.3$). The ALEQ demonstrated good validity in past research (Hankin, 2008; Hankin, Stone, & Wright, 2010), including significant correlations with ratings of episodic stressors ($r = .44$, $p < .001$) from a contextual stress interview (Rudolph & Flynn, 2007). In sum, the ALEQ possesses strong psychometric properties and provides a reasonably reliable, valid, and developmentally appropriate assessment of stressors among youth.

Genotyping. Children provided buccal cells for DNA collection via Oragene[®] kits from DNA Genotek (Ottawa, Ontario, Canada), and genomic DNA was collected and isolated using standard salting out and solvent precipitation methods. The 5-HTTLPR alleles were assayed (Anchordoquy, McGeary, Liu, Krauter, & Smolen, 2003) and modified by using primers reported by Hu et al. (2005). The rs25531 SNP genotypes (L_A vs. L_G) were obtained by incubating the PCR products with MspI (Wendland, Martin, Kruse, Lesch, & Murphy, 2006). Samples were analyzed on an ABI PRISM[®] 3130xl Sequencer. For the triallelic analyses, three groups of participants were formed based on their genotyping: children homozygous for the lower expressing S or L_G alleles, those heterozygous, and those homozygous for the higher expressing L_A allele. For biallelic analyses, trichotomous groups of SS, SL, and LL genotypes were formed.

RESULTS

Preliminary Analyses

Table 1 shows descriptive statistics, as an average across time and across informants, for the whole sample and by

grade and sex. Sex differences were observed for MASC. Specifically, girls on average across time reported more anxious symptoms compared with boys. Grade differences were noted for CDI and ALEQ. Youth in the ninth grade reported more depressive symptoms and stressors on average over time. Depressive and anxious symptoms were moderately correlated across time points (r s = .26 at baseline, .36 at Time 2, .25 at Time 3, .13 at Time 4, and .27 at Time 5). 5-HTTLPR polymorphisms were in Hardy-Weinberg equilibrium. Genotype frequencies for biallelic 5-HTTLPR were 33% L/L, 46% S/L, and 21% S/S. Genotype frequencies for triallelic 5-HTTLPR were 21% L_A homozygotes, 46% heterozygotes, and 33% homozygotes for S/ L_G . Neither genotype nor allelic frequencies varied significantly by race (White vs. African American; $\chi^2 < 1.37$).

Statistical Approach

Hierarchical linear modeling (HLM 5.04; Raudenbush, Bryk, Cheong, & Congdon, 2001) was used to investigate the main study questions. We used lagged analyses to test for 5-HTTLPR interacting with stressors assessed over time to predict *prospective elevations* in symptoms specifically. Symptoms (depressive or anxious) at time T served as the dependent variable in the HLM analysis and time T-1 symptoms were included in Level 1 so that prospective increases in symptoms from one wave to the next were predicted after controlling for prior wave's symptoms. To test whether 5-HTTLPR interacts with *idiographic* stressors, we centered ALEQ scores around each youth's mean level over time (group-mean centering). This analysis tests the stress reactivity hypothesis that 5-HTTLPR moderates the longitudinal association between increases in stressors, relative to the individual's average level, and prospective elevations in symptoms over time. In contrast, we tested the *nomothetic* approach by leaving stress scores uncentered. We controlled for age, race, and sex in all analyses.

TABLE 1
Descriptive Statistics Overall and by Sex and Grade Group

Variable	Full Sample M (SD)	Girls M (SD)	Boys M (SD)	3rd Grade M (SD)	6th Grade M (SD)	9th Grade M (SD)
CDI	6.47 (5.20)	7.17 (5.28)	5.98 (5.64)	6.19 (5.28)	6.17 (4.77)	7.28 (5.72)*
MASC	43.01 (19.89)	46.22 (21.26)	38.58 (18.35)***	43.43 (19.04)	41.26 (13.84)	40.69 (15.36)
ALEQ	16.02 (7.85)	16.41 (7.66)	15.48 (8.12)	12.96 (7.26)	15.12 (7.72)	21.02 (6.72)***

Note. All measures are composites of parent and child report and are averaged across the five time points. CDI = Children's Depression Inventory; MASC = Multidimensional Anxiety Scale for Children; ALEQ = Adolescent Life Events Questionnaires.

* $p < .05$. *** $p < .001$.

5-HTTLPR Interacts with Idiographic Stressors to Predict Prospective Elevations in Depressive Symptoms Specifically

We first examined potential gene-environment correlations (r_{GE}). There was no significant association between 5-HTTLPR and stressors at any time point ($r_s < .08$, ns). However, high-risk 5-HTTLPR alleles, when analyzed in the biallelic manner with three groups (S/S, S/L, and L/L), interacted with greater idiographic stress exposure to predict prospective elevations in depressive symptoms over 1 year (top portion of Table 2). Following up this significant interaction to examine the effect of stressors on depressive for each genotype group showed that youth with two S copies exhibited the largest association between stressors and depressive symptoms ($b = .15$, $SE = .03$, $t = 3.98$, $p < .001$; $r_{\text{effect size}} = .43$), followed by youth with S/L

TABLE 2
5-HTTLPR \times Stress Interactions Predicting Prospective Elevations of Depressive Symptoms Over 1 Year

Predictor	<i>b</i>	<i>SE</i>	<i>t</i>	<i>df</i>	<i>ES(r)</i>
<i>Idiographic Stress Reactivity</i>					
Fixed Effects					
Level 1					
CDI T-1	.65	.02	25.69***	1,219	.87
Idiographic Stress	.10	.02	4.15***	1,218	.27
Level 2					
Race	-.66	.85	-0.78	1,216	.05
Sex	.38	.23	1.62	1,216	.11
SERT	.83	.57	1.44	1,216	.10
SERT \times Idiographic Stress	.08	.02	3.79***	1,218	.25
Variance components					
Level 1 Within Person	3.36				
Level 2 Initial Status	8.67***				
Stress Slope	.004***				
CDI T-1	.068***				
<i>Nomothetic Stress Reactivity</i>					
Fixed Effects					
Level 1					
CDI T-1	.61	.03	22.68***	1,219	.83
Nomothetic Stress	.09	.02	3.95***	1,218	.26
Level 2					
Race	-.34	.35	.98	1,216	.06
Sex	.47	.51	.92	1,216	.06
SERT	.30	.36	.82	1,216	.05
SERT \times Nomothetic Stress	.01	.02	.88	1,218	.05
Variance Components					
Level 1 Within Person	3.33				
Level 2 Initial Status	13.75***				
Stress Slope	.004***				
CDI T-1	.068***				

Note. Generally, $ES(r)$.10-.23 is viewed as small, moderate is .24-.36, and greater than .37 is large, respectively (Cohen, 1988). CDI = Children's Depression Inventory; CDI T-1 = prior wave CDI score; SERT = serotonin transporter promoter polymorphism; $ES(r)$ = effect size r (Rosnow & Rosenthal, 1991).

* $p < .05$. ** $p < .01$. *** $p < .001$.

($b = .09$, $SE = .02$, $t = 3.34$, $p < .001$; $r_{\text{effect size}} = .37$), and last by homozygous L carriers ($b = .05$, $SE = .03$, $t = 1.37$, $p = .17$; $r_{\text{effect size}} = .16$).

This $G \times E$ effect is shown in Figure 1 with stressors depicted at 1 SD above and below the mean for idiographic ALEQ scores. Youth with two short copies of 5-HTTLPR (S/S) who experienced more stressors over time, relative to their typical exposure to stressors, reported the greatest prospective increases in depressive symptoms. In contrast, when stressors were analyzed in a nomothetic manner, the $G \times E$ effect was not significant (bottom portion of Table 2). The main effect of nomothetic stressors predicted prospective elevations in depressive symptoms, consistent with Risch and colleagues' meta-analysis (2009). Moreover, when 5-HTTLPR was genotyped using the triallelic genotyping with trichotomous grouping, similar results were observed as the $G \times E$ ($b = .06$, $SE = .02$, $t = 2.86$, $p < .01$, $r_{\text{effect size}} = .19$) was significant in predicting prospective elevations in depressive symptoms. Finally, neither sex ($b = .005$, $SE = .004$, $t = 1.36$, $p = .17$, $r_{\text{effect size}} = .09$) nor grade ($b = .023$, $SE = .015$, $t = 1.54$, $p = .09$, $r_{\text{effect size}} = .10$) moderated the $G \times E$ effect.

Specificity of Effects

This specificity analysis examines the unique prediction of $G \times E$ for depressive versus anxious symptoms given their known co-occurrence. Similar lagged HLM analyses were conducted as previously described except that composite MASC anxiety scores served as the dependent variable. We entered MASC scores at T-1 and idiographic stressors at Level 1, 5-HTTLPR at Level 2, and the cross-level interaction of 5-HTTLPR \times Idiographic Stressors. Results showed main effects for biallelic genotyping of 5-HTTLPR ($b = 2.86$, $SE = 1.12$, $t = 2.38$,

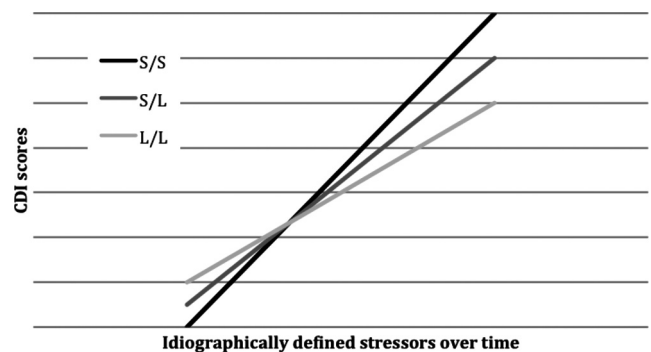


FIGURE 1 Interaction between 5-HTTLPR and idiographic stressors (combined parent and youth report) predicts prospective elevations in depressive symptoms (combined parent and youth report) over time. S = short; L = long; CDI = Children's Depression Inventory.

$p < .05$, $r_{\text{effect size}} = .19$), prior anxiety symptoms ($b = 1.98$, $SE = .42$, $t = 4.75$, $p < .001$, $r_{\text{effect size}} = .30$), and idiographic stressors ($b = .88$, $SE = .09$, $t = 9.96$, $p < .001$, $r_{\text{effect size}} = .56$), yet no significant interaction ($b = .14$, $SE = .12$, $t = 1.17$, $p = .24$, $r_{\text{effect size}} = .08$). Likewise, no significant interaction was observed for the nomothetic stress approach ($b = .14$, $SE = .18$, $t = .67$, $p = .51$, $r_{\text{effect size}} = .04$). Similar results were observed when triallelic genotyping trichotomous groups were analyzed. Such nonsignificant G × E for anxiety symptoms demonstrates specificity to depression.

Finally, a more stringent test of specificity and sensitivity is to examine whether 5-HTTLPR × Idiographic Stressors continues to predict prospective elevations in depressive symptoms over time after controlling for the concurrent co-occurrence of anxious symptoms. In this final analysis, we used the same HLM equation for 5-HTTLPR × Idiographic Stress predicting prospective elevations in depressive symptoms except that MASC scores at Time T were also included at Level 1 to control for the overlap between anxious and depressive symptoms. The biallelic genotyped 5-HTTLPR × Idiographic Stress interaction still significantly predicted prospective elevations in depressive symptoms ($b = .087$, $SE = .02$, $t = 4.14$, $p < .001$, $r_{\text{effect size}} = .26$), and the within person main effect of MASC was likewise significant ($b = .045$, $SE = .008$, $t = 5.69$, $p < .001$, $r_{\text{effect size}} = .36$). The same pattern was seen with triallelic genotyping. In sum, youth carrying high-risk susceptibility 5-HTTLPR alleles who experienced greater idiographic stress exposure experienced increases in prospective elevations in depressive symptoms over 1 year even after accounting for the co-occurring influence of anxiety symptoms.

DISCUSSION

Theoretical and empirical interest in G × E has surged since Caspi and colleagues' seminal findings, yet there has been doubt concerning the replicability of G × E in depression since Risch et al.'s (2009) meta-analysis. Also, the extent to which measured genes interact with stressors among youth to predict prospective elevations in depressive symptoms specifically has been unclear. Results from this study indicate that youth who carry at least one S allele of the variant of 5-HTTLPR exhibited higher prospective elevations in depressive symptoms specifically over a 1-year follow-up when they experienced more stressors relative to their own typical exposure to stress. This overall significant G × E effect represented a moderate effect size. This 5-HTTLPR × Idiographic Stressor interaction is consistent with a stress reactivity conceptualization of how 5-HTTLPR may enhance susceptibility to depression in concert with

environmental risk (cf. Caspi et al., 2010). No G × E effect was observed when stressors were analyzed in a nomothetic manner. These findings were observed for both boys and girls as well as youth of all ages as neither sex nor grade moderated the G × E effect. Last, this G × E effect specifically predicted depressive, not anxious, symptoms. In sum, this study provides evidence for G × E effects in youth depression using a multiwave longitudinal design with reliable, developmentally valid assessments of proximal stressors (Monroe & Reid, 2008).

An exciting and unique contribution of the present study is the use of a multiwave design to enable an *idiographic* as well as a *nomothetic* approach (Abela & Hankin, 2008) to analyze stress reactivity and environmental risk in G × E. The significant G × E effect predicting prospective elevations in depressive symptoms was observed only when analyzed from an idiographic conceptualization, in which individuals' increases or decreases in stressors over time were compared relative to their own average exposure to stress. We based our a priori prediction that G × E effects would hold when conceptualized and analyzed from an idiographic perspective due to consistent and robust results for stress reactivity demonstrated in cognitive vulnerability-stress research using an idiographic versus nomothetic approach (Abela & Hankin, 2008). This study's findings suggest that 5-HTTLPR allelic variation may underlie individuals' stress reactivity to increases in one's typical exposure to stressors, as indicated by the significant genetic interaction with idiographic stressors. The present study used a multiwave design in G × E depression research that provides a strong test of whether 5-HTTLPR conveys susceptibility to depression, in concert with idiographic stress exposure, via stress reactivity processes. This differentiation of environmental stress conceptualization and analysis (i.e., idiographic and nomothetic) offers another potential explanation for the lack of an overall significant G × E effect in the Risch et al. (2009) meta-analysis (see Caspi et al., 2010; Rutter et al., 2009; Uher & McGuffin, 2010, for other limitations), as stress reactivity has been conceptualized and analyzed in a purely nomothetical manner in all prior G × E research. Taking the present evidence that genetically susceptible (at least 5-HTTLPR) youth who are exposed to idiographic stressors exhibited the greatest elevations in depressive symptoms in concert with prior human (e.g., Gotlib et al., 2008; Hariri et al., 2002) and animal (e.g., Li et al., 1999) research, an increasing and coherent literature is building suggesting that 5-HTTLPR may confer risk to depression via stress reactivity processes (cf. Caspi et al., 2010).

A few prior studies have found a G × E predicting depressive symptoms specifically in females (e.g., Eley et al., 2004; Hammen et al., 2010), yet our data did

not support sex moderation for the $G \times E$ effect. The precise reason for the inconsistencies across studies is unclear, although none of the previous studies finding sex moderation examined environmental stress from an idiographic conceptualization. In accordance with our findings, studies investigating the role of stress reactivity in relation to depression have found this effect across both male and female individuals (Hankin, Badanes, Abela, & Watamura, 2010; Hariri et al., 2002; Li et al., 1999; Williams et al., 2009). Our evidence suggests a plausible psychobiological process for how the $G \times E$ effect in depression works in both male and female individuals and moves beyond merely investigating $G \times E$ interactions with purely statistical techniques by examining $G \times E$ effects using psychological and biologically plausible approaches and conceptualizations (Rutter et al., 2009). Finally, grade cohort did not moderate the significant $G \times E$ effect in this sample. As the majority of youth $G \times E$ research has used adolescent samples and little research has investigated $G \times E$ in prepubertal children, the lack of moderation by grade cohort suggests that 5-HTTLPR may confer susceptibility to depressive symptoms when both children and adolescents experience idiographic stressors. Clearly, though, additional research with younger children is needed to replicate this finding.

Of interest, when our data were analyzed in an idiographic manner, a slight cross-over interaction was observed (see Figure 1). These types of interactions have also been observed in previous $G \times E$ studies predicting depression (e.g., Caspi et al., 2003; Eley et al., 2004; Wilhelm et al., 2006). The implications of these findings have been extensively discussed in relation to Belsky's (1997) differential-susceptibility hypothesis and Boyce and Ellis's (2005) biological-sensitivity-to-context thesis (also see Belsky & Pluess, 2009a, 2009b; Ellis & Boyce, 2008). Both of these theoretical arguments emphasize the importance of examining how certain risk factors, such as the S allele of the 5-HTTLPR gene, could be a marker of environmental sensitivity in which those possessing one or both of these alleles would either thrive in an enriched environment or be especially vulnerable in an adverse environment (Belsky & Pluess, 2009a). As seen in Figure 1, our data demonstrate that the participants who carry either one or two copies of the S allele tend to have lower levels of depressive symptoms in the presence of low idiographic stress, as well as higher levels of depressive symptoms in the presence of high idiographic stress when compared to those homozygous for the long allele. Our findings lend support to both Belsky and Boyce and Ellis's conceptualization of how S 5-HTTLPR alleles may be related to differential-susceptibility to environment as opposed to the conceptualization that possessing an S allele confers only risk or vulnerability.

Strengths and Limitations

Several features of the present study are unique and provide a more methodologically and conceptually rigorous test of hypotheses that enhance confidence in the results and advance knowledge of $G \times E$ in youth depression. The study included a multiwave, multi-informant design with a community-based sample, which increased statistical power and enabled us to evaluate both idiographic and nomothetic approaches to measuring stress. In addition, we combined child and parent reports of youths' anxious and depressive symptoms and stressors to reduce reliance on a single source. We used a general community sample of youth, as opposed to a clinical sample with biases that are known to reduce generalizability and accurate statistical tests (Cohen & Cohen, 1984).

Next, we examined $G \times E$ specificity to depression versus anxiety given the strong co-occurrence of these internalizing symptoms and found unique prediction to depression. This finding provides some important information about specific etiological processes predicting depression versus anxiety. Of note, our findings supported a main effect of the 5-HTTLPR gene for anxiety symptoms. Meta-analyses have not found support for the 5-HTTLPR polymorphism predicting certain subtypes of anxiety in adults, such as harm avoidance, but have found a main effect predicting neuroticism (Munafò et al., 2008; Schinka, Busch, & Robichaux-Keene, 2004). Indeed, research examining main effects of 5-HTTLPR on anxiety has been relatively rare (Gonda et al., 2009), and finding this direct association, especially with youth, is novel.

Last, we used both biallelic and triallelic method for genotyping 5-HTTLPR (Hu et al., 2006; Hu et al., 2005). $G \times E$ effects were found for both approaches. We used relatively reliable, valid, and developmentally appropriate assessments of negative life events, as reported by both youth and parent, to capture stress exposure as environmental risk to test $G \times E$ more accurately.

Still, study limitations provide avenues for future research. First, the sample size was relatively modest, although significant $G \times E$ was observed when the appropriate characterization of environmental stress (idiographic) was used. Also mitigating the concern of sample size, reviewers of the $G \times E$ literature (e.g., Uher & McGuffin, 2008, 2010) suggest minimum sample sizes, based on power considerations (e.g., Luan, Wong, Day, & Wareham, 2001), and the present study exceeded minimum thresholds, at least those recommended for cross-sectional $G \times E$ research. Still, there is the possibility that the present findings may represent chance effects, and replication of these effects are needed with larger samples that assess stressors, preferably with

reliable and valid methods, repeatedly over time. Alternatively, it may be that relatively modest instead of significantly more sizeable sample sizes are sufficient to detect $G \times E$, especially when investigated with a multiwave design, which increases statistical power (Snijders & Bosker, 1999); accurate measurement of genes and environment (Monroe & Reid, 2008), which reduce measurement error; and a construct-validity, theory-driven approach (Caspi et al., 2010). Indeed, prior work has demonstrated that inaccurate assessments decrease statistical power to such a substantial degree that smaller samples with accurate measurement may be preferable (Wong, Day, Luan, Chan, & Wareham, 2003).

Second, the significant idiographic $G \times E$ predicted prospective elevations in depressive symptoms, but we did not assess clinical depression. Most research suggests that depression can be represented and conceptualized best as a dimensional continuum rather than discrete category (Hankin, Fraley, Lahey, & Waldman, 2005). Also, subclinical depressive symptoms are important to study in their own right as they represent more than mere “moodiness” and predict diagnosable disorder (Klein, Shankman, Lewinsohn, & Seeley, 2009; Pine, Cohen, Cohen, & Brook, 1999) and other psychosocial impairments (Gotlib, Lewinsohn, & Seeley, 1995). The percentage of youth with depressive symptoms scores above normative cutoffs was within the range expected in unselected community samples of youth of these ages (Cole et al., 1998; Petersen et al., 1993). Thus, although the current sample reported relatively low levels of depressive symptoms, understanding $G \times E$ effects predicting elevations in subclinical depressive symptoms over time, especially among youth, is an important area of research. Nevertheless, diagnostic interviews to ascertain clinical depression across multiple assessments can address potential concerns about continuity of findings to depressive disorder. In addition, the majority of our sample was comprised of younger youth (72% in the third or sixth grade), and research has demonstrated that significant increases in depression generally occur after age 13 (Hankin et al., 1998). Although we controlled for age in our analyses and it can be considered a strength that our sample was mostly children and early adolescents, future research is still needed to investigate possible developmental influences in $G \times E$ processes over time and whether static genetic polymorphisms interact with increases in stressors during adolescence to predict the surge in depression in middle to late adolescence.

Third, some have criticized self-report of stressors and argued that questionnaire assessments may not be as rigorous or accurate compared with contextual stress interviews (Monroe & Reid, 2008). The stress measure used in this study has shown good reliability and validity

in prior research (e.g., Hankin, 2008; Hankin, Stone, et al., 2010), including significant and moderate correlations with a contextual stress interview. Also, the stress measure uses very specific language with precise anchors and examples to define stressors as recommended to more accurately assess negative events (Brewin et al., 1993; Dohrenwend, 2006). We assessed stressor occurrence with multiple informants (parent and youth) to reduce subjectivity unreliability and overreporting of trivial events. These methodological features match those used by Caspi and colleagues (2003). Still, the use of a self-reported measure of negative life events is a limitation of the study. Future $G \times E$ research examining stress would benefit from use of contextual stress interviews, as seen in research by Hammen and colleagues (2010), and doing so in a multiwave design.

Finally, we analyzed DNA sequence variation of one specific allelic variation in 5-HTTLPR. The environment has a significant impact on gene expression and outcomes through epigenetic processes (Tsankova, Renthal, Kumar, & Nestler, 2007). Incorporating epigenetics (Mill & Petronis, 2007) and static DNA sequence variation of other genes in addition to 5-HTTLPR is a clear priority for future research. For example, Meaney and colleagues' research with mice (Weaver et al., 2004) and human postmortem brain tissue (McGowan et al., 2009) reveals epigenetic influences of the glucocorticoid receptor in response to environmental influences. Similar epigenetic effects accounting for $G \times E$ is likely in the development of psychopathology (Mill & Petronis, 2007; Tsankova et al., 2007).

Implications for Research, Policy, and Practice

Understanding what risk factors and processes predict increases in depressive symptoms among youth is clearly important for policy and practice. Elevated depressive symptoms are among the most potent predictors of later onset of clinical depression (Klein et al., 2009). Youth with subthreshold depression represent important targets for selective interventions to forestall onset of major depression and subsequent interpersonal and academic difficulties associated with elevated depression (Gotlib et al., 1995). Most recently, Garber and colleagues (2009) demonstrated that onset of major depression could be significantly prevented and delayed among high-risk youth, as defined by elevated, subthreshold levels of depression and/or parents with current or past depression. As the present study showed that genetic risk, specifically 5-HTTLPR, interacts with idiographic increases in stressors to prospectively predict elevations in depressive symptoms, future preventive research may benefit by incorporating an integrated genetic perspective to further understand how and for whom selective

interventions and prevention efforts work to reduce depression onset (cf. Uher, 2008).

In sum, this study showed that youth who experience more idiographic stressors over time are most likely to exhibit prospective increases in depressive symptoms specifically, especially when those youth are at measured genetic risk. These data suggest that 5-HTTLPR may confer susceptibility to depression through the process of stress reactivity in that youth carrying at least one copy of the 5-HTTLPR lower expressing variant may be more susceptible to react to their environment (Belsky, Bakermans-Kranenburg, & van IJzendoorn, 2007; Caspi et al., 2010). When that environment is characterized by relatively greater exposure to stressful negative life events, youth with 5-HTTLPR susceptibility are more likely to experience increases in depressive symptoms specifically over time.

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